



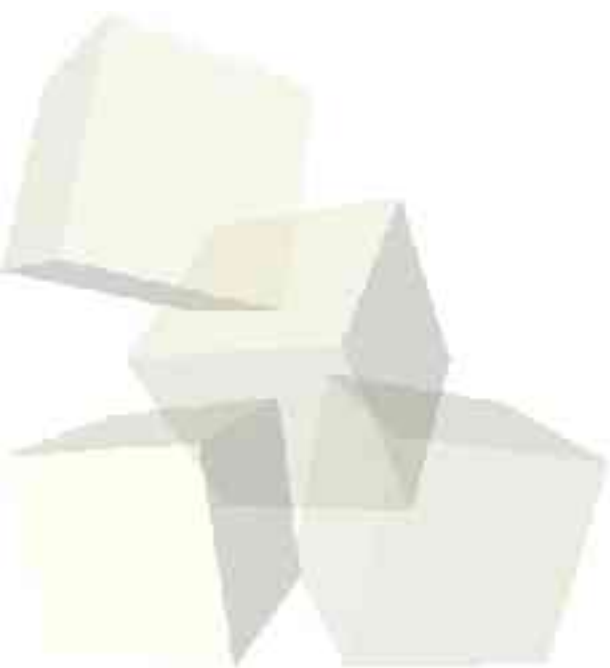
Time Dependent Update Functions for Perfect Sampling

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A simple logic puzzle

Results for the Putnam Mathematical Competition in 2000

Harvard

MIT

Duke

Caltech

	1	2	3	4
Harvard	Orange	Orange	Pink	Blue
MIT	Orange	Orange	Orange	Pink
Duke	Pink	Blue	Blue	Blue
Caltech	Orange	Pink	Blue	Blue

Facts:

- Harvard came in 3 (or better)
- Caltech came in 2 (or better)
- Duke came in first

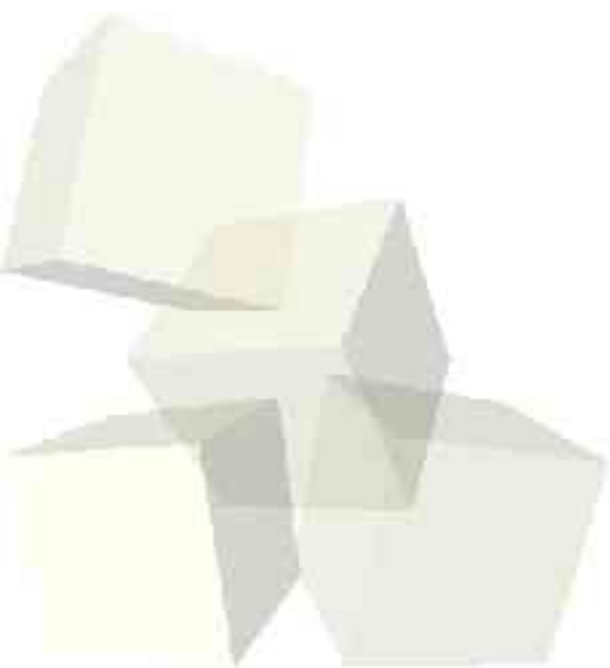


Simple idea can form the basis for algorithms for generating weighted permutations

Each permutation x given weight $\mu(x)$

Goal generate random variates from

$$\pi(x) = \frac{\mu(x)}{\sum_y \mu(y)}$$





- A weighted permutation problem
- Perfect sampling with CFTP
- Time dependent update functions
- Bounding chains
- BC for weighted permutations
 - does not work with original CFTP
- Solving the permutation problem
- A continuous example

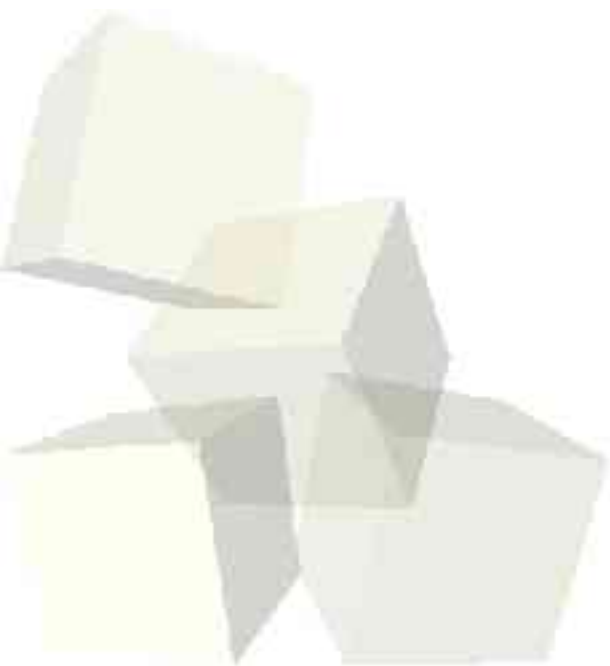


A permutation problem

The Goal

Generate uniformly from the set of random linear extensions of a poset

$$\mu(x) = \begin{cases} 1 & \text{if } x \text{ is a linear extension} \\ 0 & \text{otherwise} \end{cases}$$





Linear Extensions

Let $N = \{1, 2, \dots, n\}$

A partial order \leq on N is

1) **Reflexive** $a \leq a$

2) **Antisymmetric** $a \leq b$ and $b \leq a$ implies $a = b$

3) **Transitive** $a \leq b$ and $b \leq c$ implies $a \leq c$

A linear extension is a permutation x where

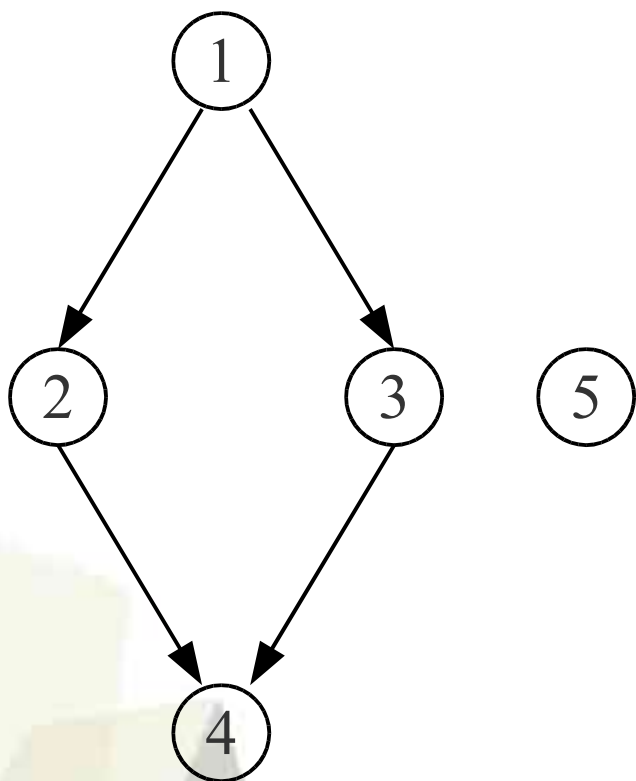
$$a \leq b \text{ implies } x(a) \leq x(b)$$

Suppose throughout $1, 2, \dots, n$ is a linear ext.



Example

$n=5$



Say item 3 is in position 2

Some Linear Extensions

1 2 3 4 5

1 3 2 4 5

1 2 3 5 4

1 3 2 5 4

1 2 5 3 4

1 3 5 2 4

⋮



Question: How many linear extensions?

Bad news: #P complete

[Brightwell, Winkler 1991]

Good news: selfreducible so generating samples leads to efficient (approx) counting

[Jerrum, Valiant, Vazirani 1986]

Question: Average # of inversions?

Variant of nonparametric Kendall's tau test

[Efron, Petrosian 1999]



Markov chain approach

Describe Markov chain via update function

$$f : \Omega \times [0,1] \rightarrow \Omega$$

$$U_1, U_2, \dots \sim \text{Unif}[0,1] \quad (\text{iid})$$

$$X_{t+1} = f(X_t, U_{t+1})$$

Computer simulation of Markov chains
(use pseudorandom numbers)

Also known as

transition function

stochastic recursive scheme



Markov chain approach

One step in linear extensions chain

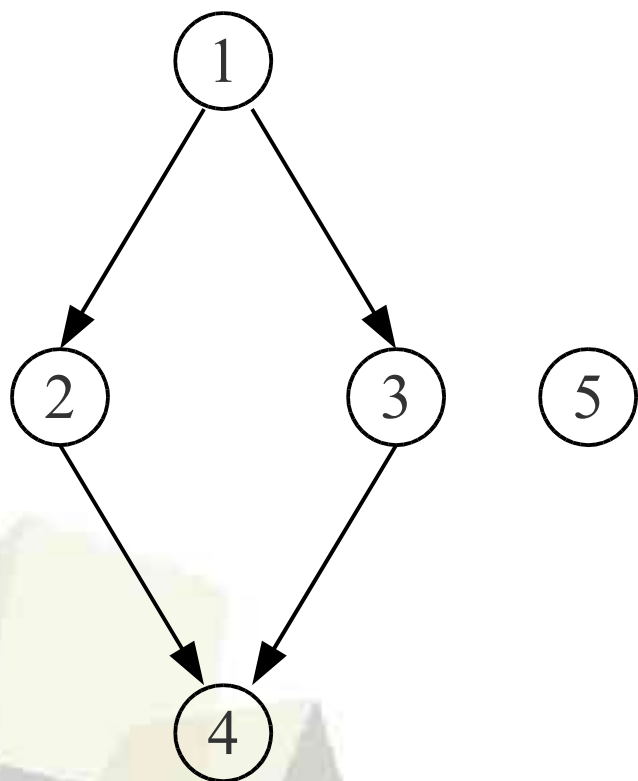
$$y = f(x, U)$$

- 1] Choose i uniformly from $\{1, 2, \dots, n-1\}$
- 2] Choose B uniformly from $\{0, 1\}$
- 3] Let $y \leftarrow x$
- 4] If $x(i) \leq x(i+1)$ or $B=0$ do nothing
Else $y(i+1) \leftarrow x(i), y(i) \leftarrow x(i+1)$

(Pick random position, flip fair coin,
if coin heads and does not violate partial order
then swap items at position and one to the right)



Example steps



State

Randomness

1 2 5 3 4

$i=2, B=1$

1 5 2 3 4

$i=3, B=1$

1 5 2 3 4

$i=1, B=0$

1 5 2 3 4



[Bubley, Dyer 1999] Different Markov chain mixes $O(n^3 \ln(n/\epsilon))$ steps

[Wilson 2004] Original chain mixes $O(n^3 \ln(n/\epsilon))$

[Felsner, Wernich 1997] Perfect sampling when poset two dimensional

This work: perfect sampling arbitrary poset

$$O(n^3 \ln(n))$$



Perfect Sampling Algorithm

A perfect sampling algorithm has 3 properties:

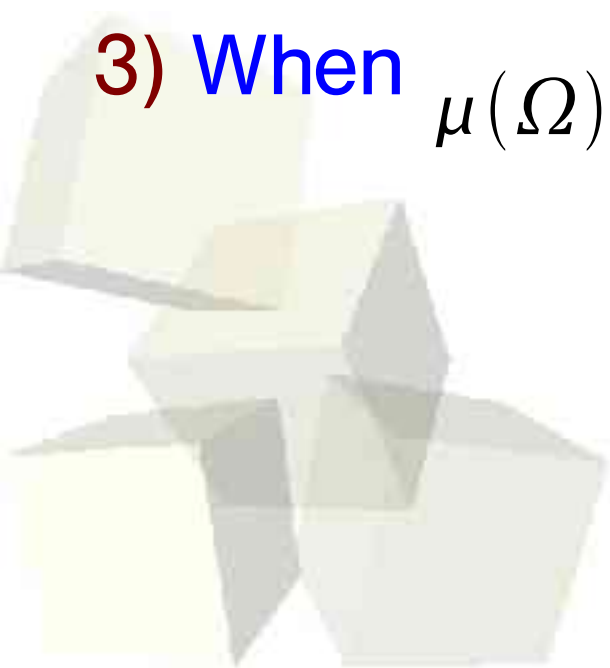
- 1) Generates exact random variates from the target distribution
- 2) Running time is random with exponential tails

$$P(T > 2kE[T]) < (1/2)^k$$

- 3) No knowledge of the normalizing constant
(good since we do not know $\mu(\Omega)$)



- 1) Usually samples are only approximately from target distribution
- 2) Running time very unlikely to be much larger than expected value (otherwise Dyer showed any approximate sampler is a perfect sampler)
- 3) When $\mu(\Omega)$ known it is a direct sampler





The Good News

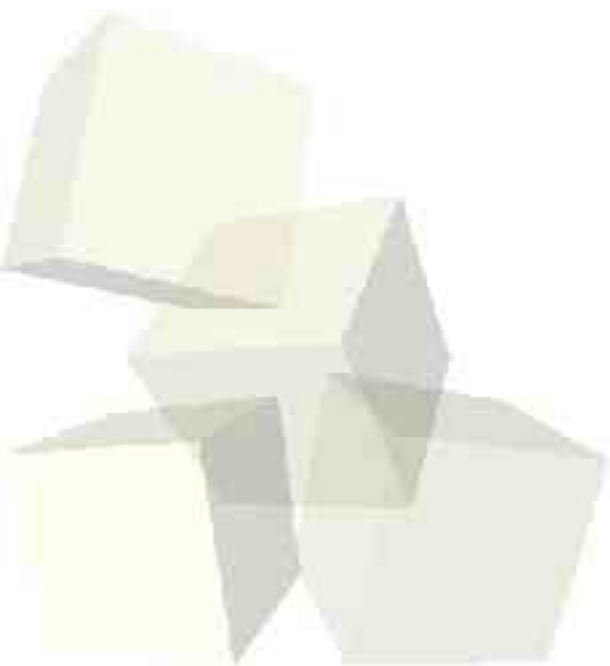
- Generates exactly from desired distribution
- Can be used for continuous or discrete
- True algorithms
(Markov chain methods are not algorithms unless the mixing time is known)
- Useful even if running time unknown





The Bad News

- Not a magic solution to slow Markov chains
- Requires more effort than Metropolis-Hastings
- Methods more complex





Coupling From the Past

CFTP [Propp, Wilson '96]

Ingredients: update function for chain,

$\dots, U_{-2}, U_{-1}, U_0 \sim \text{Unif}[0,1]$ (iid)

- 1) pretend have unknown stationary rand. var.
- 2) run chain forward fixed number of steps
- 3) if state becomes known, output
- 4) else call CFTP recursively
- 5) run chain forward fixed number of steps





An example

Example

Start $T = -5$

If had

then

$$X_{-5} \sim \pi$$

$$X_{-4} = f(X_{-5}, U_{-4})$$

$$X_{-3} = f(X_{-4}, U_{-3})$$

$$X_{-2} = f(X_{-3}, U_{-2})$$

$$X_{-1} = f(X_{-2}, U_{-1})$$

$$X_0 = f(X_{-1}, U_0)$$

output

$$X_0 \sim \pi$$



Example

Start $T = -5$

Suppose I do not know $X_{-5} \sim \pi$, set $Z_{-5} = \Omega$

let $Z_{-4} = f(Z_{-5}, U_{-4})$ and $X_{-4} \in Z_{-4}$

$Z_{-3} = f(Z_{-4}, U_{-3})$ and $X_{-3} \in Z_{-3}$

$Z_{-2} = f(Z_{-3}, U_{-2})$ and $X_{-2} \in Z_{-2}$

$Z_{-1} = f(Z_{-2}, U_{-1})$ and $X_{-1} \in Z_{-1}$

$Z_0 = f(Z_{-1}, U_0)$ and $X_0 \in Z_0$

if $Z_0 = \{x\}$, $X_0 = x$
else



Second level of CFTP

Go back farther in time...

Start $T = -100$

Do not know $X_{-100} \sim \pi$, set $Z_{-100} = \Omega$

find

Z_{-5} using U_{-99}, \dots, U_{-5}

if $Z_{-5} = \{y\}, X_{-5} = y$

$$X_0 = f(f(f(f(f(X_{-5}, U_{-4}), U_{-3}), U_{-2}), U_{-1}), U_0)$$

else



Go back farther in time...

Start $T = -1000$

Do not know $X_{-1000} \sim \pi$, set $Z_{-1000} = \Omega$
then find

Z_{-100} using $U_{-999}, \dots, U_{-100}$

if $Z_{-100} = \{y\}, X_{-100} = y$

find X_0 using U_{-99}, \dots, U_0

else [keep going back in time until success]



How to keep track of Z_t

Monotonicity [Propp, Wilson 1996]

Multigamma coupling [Murdoch, Green 1998]

Bounding chains [Häggström, Neland 1999],
[H. 1999, 2004]

Multishift coupling [Wilson 2000]



Time dependent update functions

Original CFTP

Use same update function every time

$$X_{t+1} = f(X_t, U_{t+1})$$

(is a Markovian coupling)

Time dependent CFTP

Let update function change each step

$$X_{t+1} = f(X_t, U_{t+1}, U_t, U_{t-1}, \dots)$$

(non-Markovian coupling)



Sufficient conditions on f

Our requirement

Suppose $U \sim \text{Unif}[0,1]$, then

$$P(X_{t+1} \in A | X_t = x) = P(f(x, U, u_1, u_2, \dots) \in A)$$

for all values of

$$u_1, u_2, \dots$$

Proof that Time dependent CFTP works
essentially same as original proof



Bounding chains

Is itself a Markov chain

Bound each dimension separately

Keeps set of possible values at each coordinate

For permutations, keep interval for each item

1 5 2 3 4

bounded by

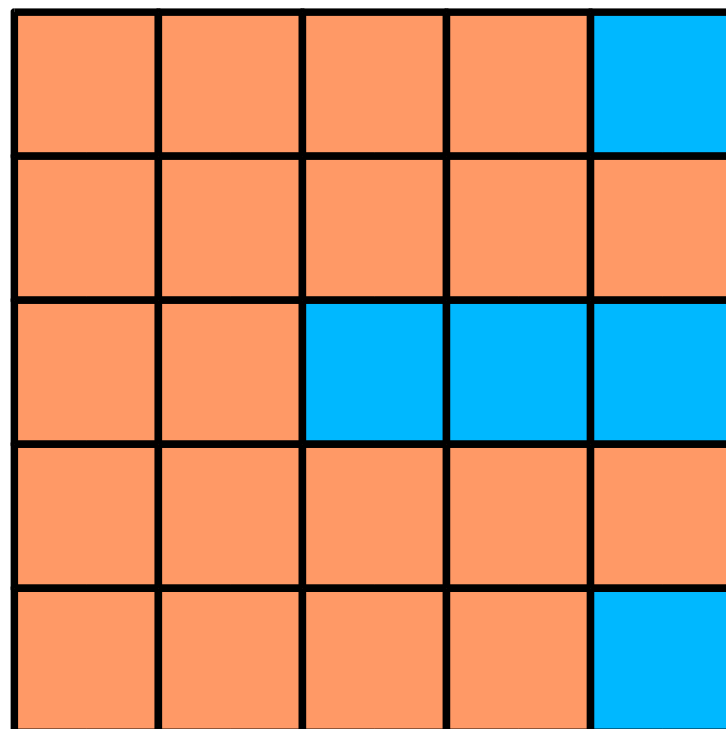
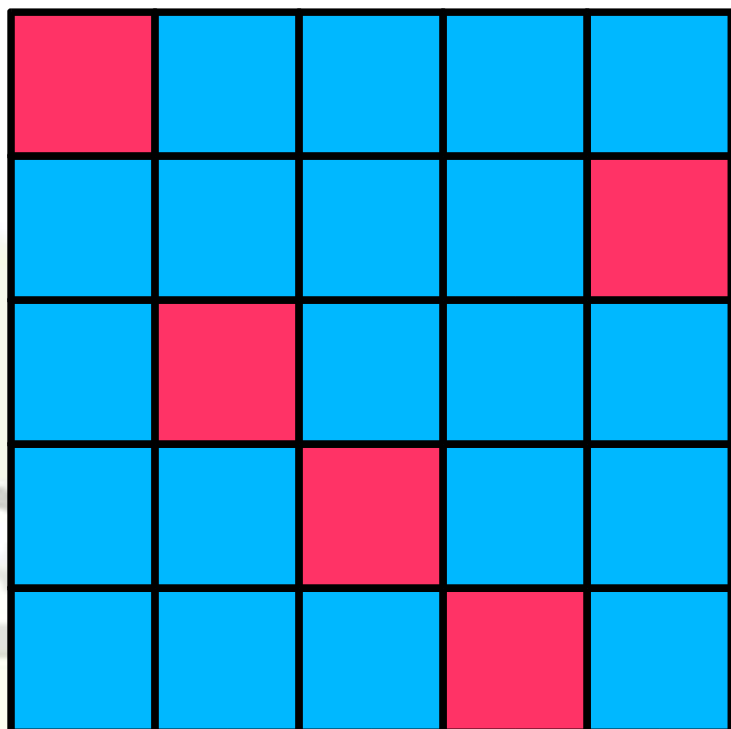
$\{1,2,3,4\}$ $\{1,2,3,4,5\}$ $\{1,2\}$ $\{1,2,3,4,5\}$ $\{1,2,3,4\}$



Pictorially

1 5 2 3 4

$\{1,2,3,4\}, \{1,2,3,4,5\},$
 $\{1,2\}, \{1,2,3,4,5\},$
 $\{1,2,3,4\}$



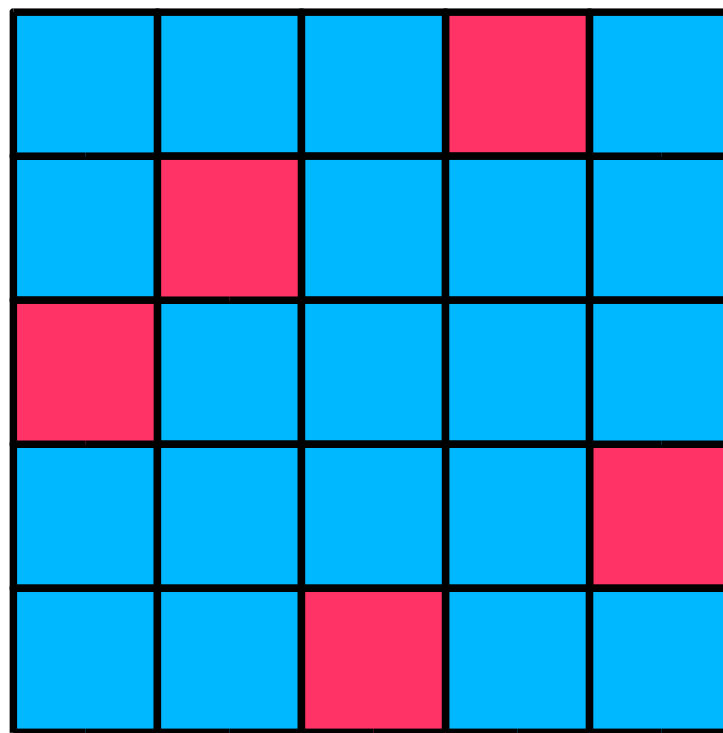
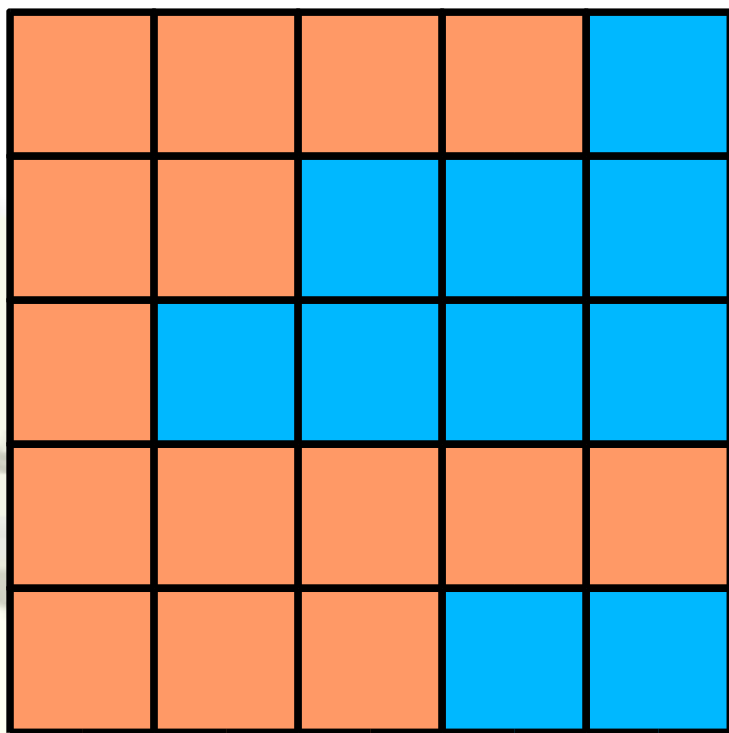


Trick from beginning

When bounding state

$$\{1, \dots, a_1\}, \dots, \{1, \dots, a_n\}$$

has a_1, \dots, a_n all different $Z_t = \{x\}$





Running the bounding chain

Suppose bounding state

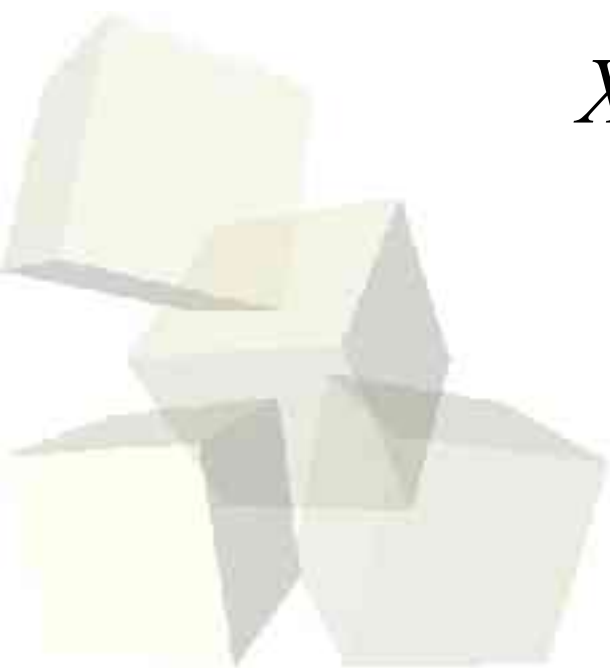
$$Y_t = \{1, \dots, a_1\}, \dots, \{1, \dots, a_n\}$$

want to run the chain forward so that if

$$X_t(i) \in \{1, \dots, a_i\} \text{ for all } i$$

then

$$X_{t+1}(i) \in \{1, \dots, a_i\} \text{ for all } i$$





Call item i active if

$$a_i \notin \{a_1, \dots, a_{i-1}, a_{i+1}, \dots, a_n\}$$

or

$$a_i = n \text{ and } a_i \notin \{a_1, \dots, a_{i-1}\}$$

position

item

orange	orange	orange	orange	blue
orange	orange	blue	blue	blue
orange	orange	orange	orange	orange
orange	orange	orange	orange	orange
orange	orange	orange	orange	orange

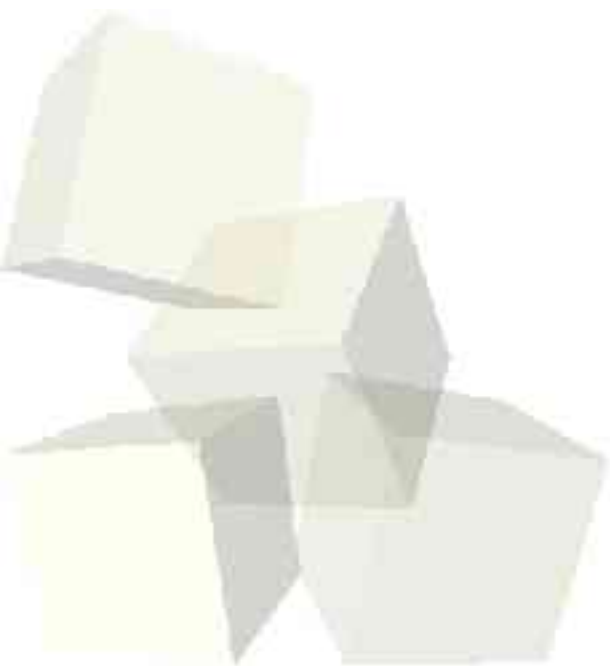
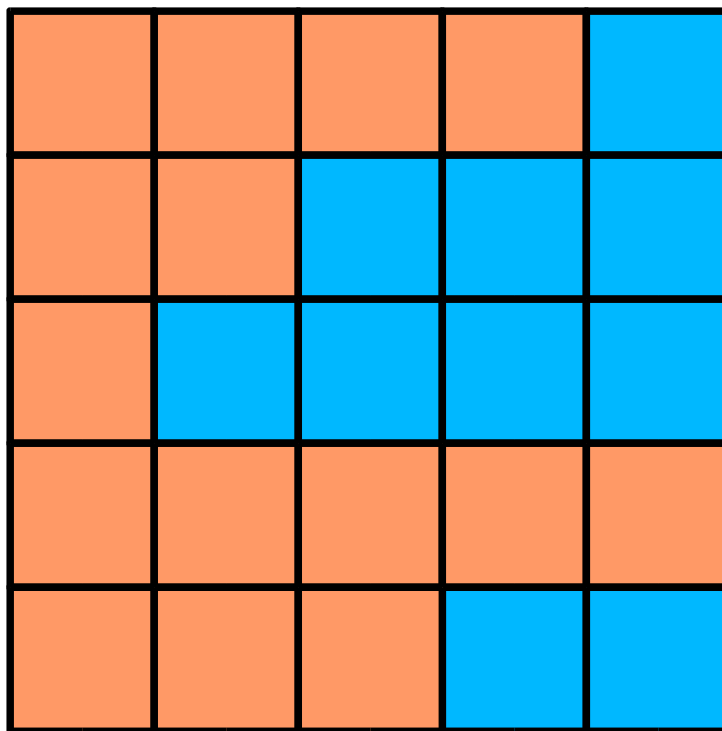
} *active items* $\{1, 2, 3\}$



Why look at active

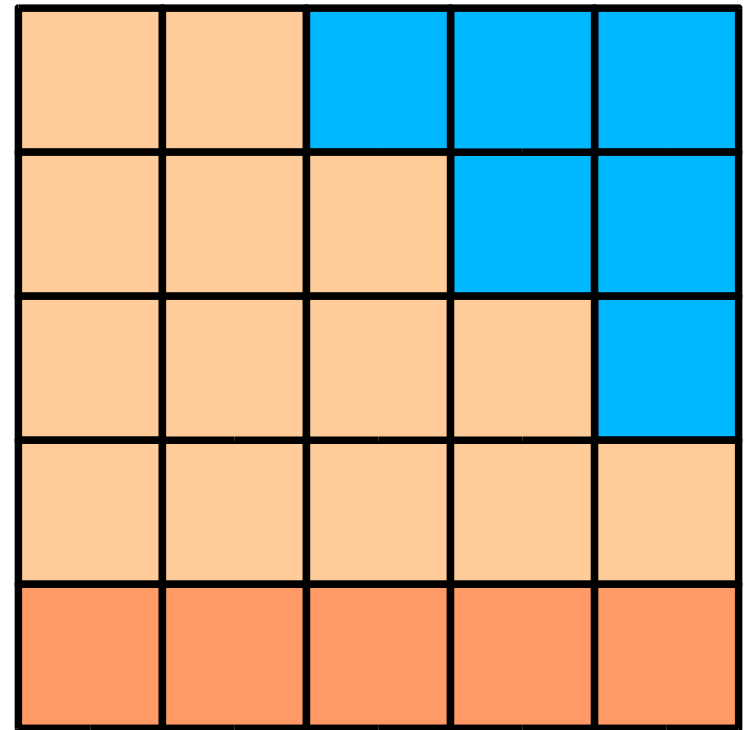
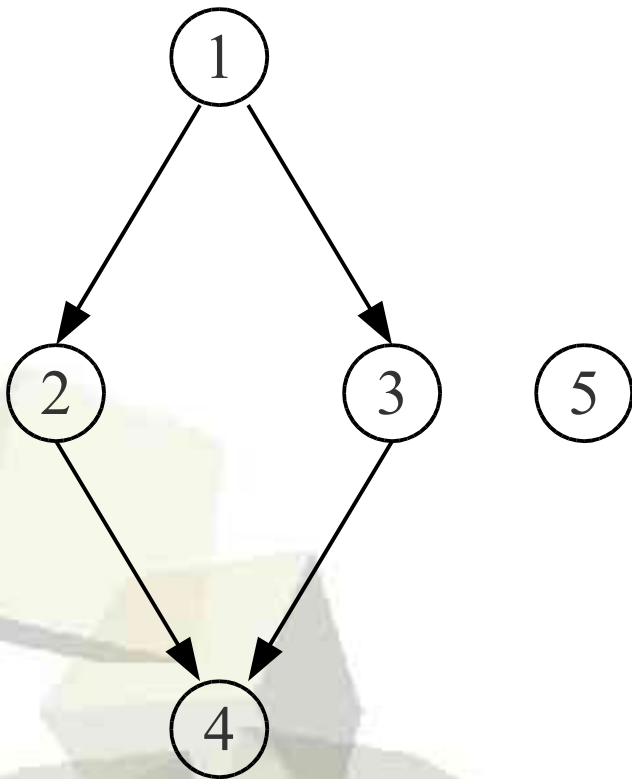
When all items active

$$Z_t = \{x\}$$



Example starting bounding chain

Since item 1 preceeds 3 items, $X_t(1) \leq 2$





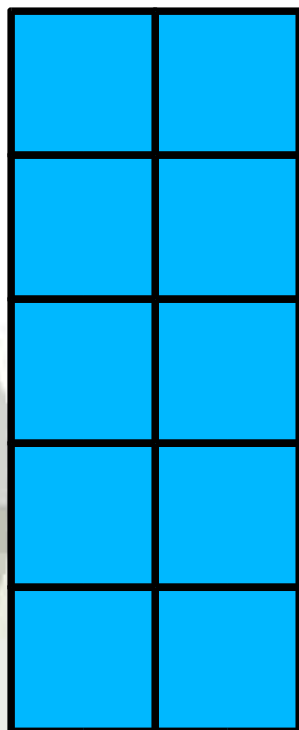
Case by case

Choose $i \sim \text{Unif}\{1, 2, \dots, n-1\}$

Choose $B \sim \text{Unif}\{0, 1\}$

Case I: no active items at position i or $i+1$

i $i+1$



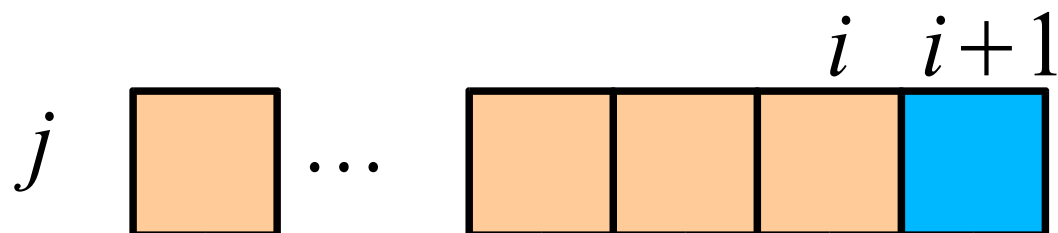
Action

bounding state unchanged

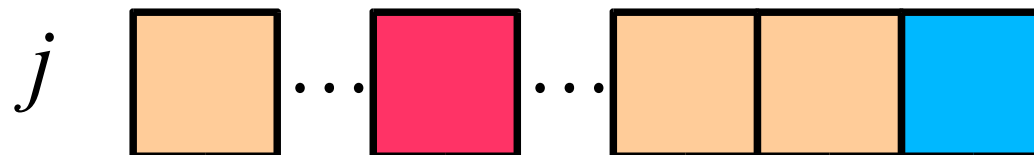
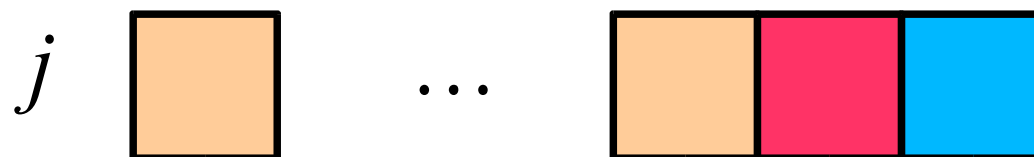
$$Y_{t+1} \leftarrow Y_t$$



Case II: one active item j at position i



Could be



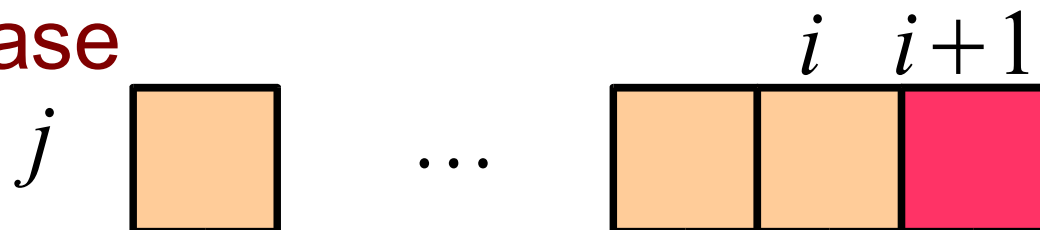
Worst case $B=1$





Case II continued

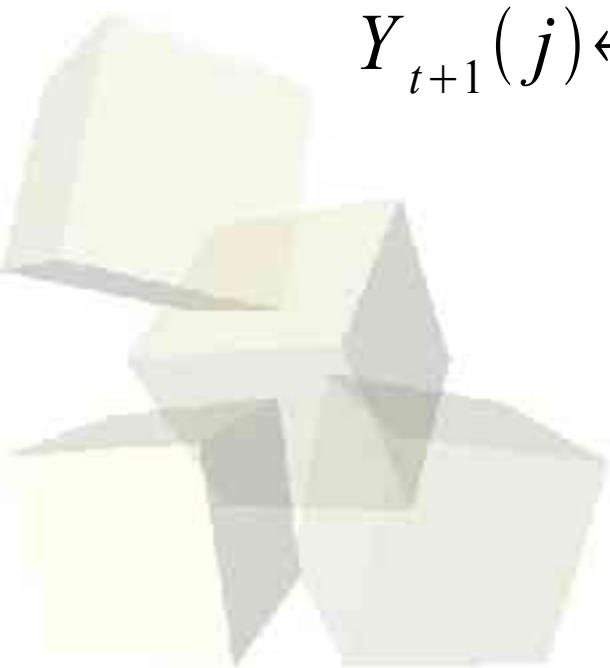
Worst case



Action

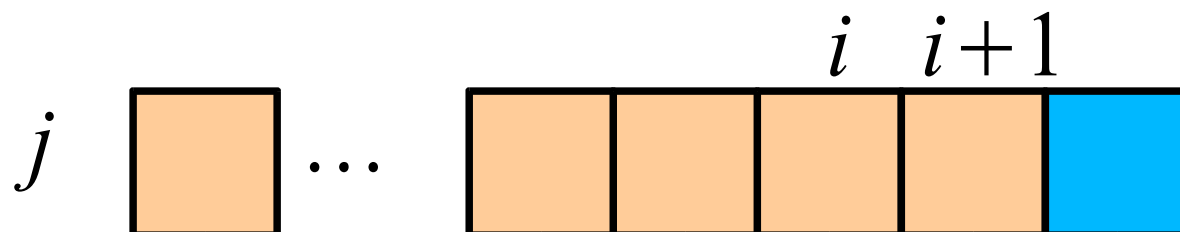
If $B=1$

$$Y_{t+1}(j) \leftarrow Y_t(j) + 1$$





Case III: one active item j at position $i+1$



Note: any other active item j' with $j' \leq j$

has

$$Y_t(j') \leq Y_t(j) - 2$$

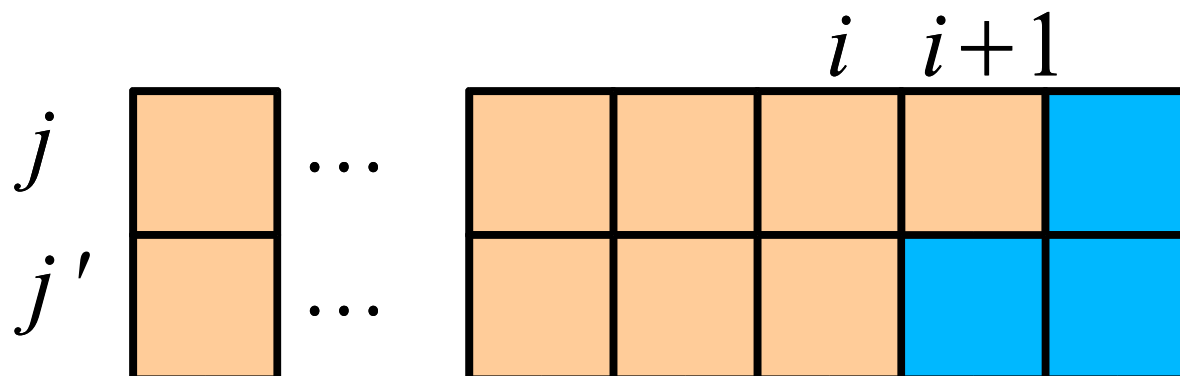
Action

If $B=1$

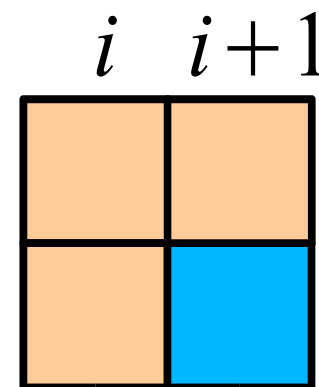
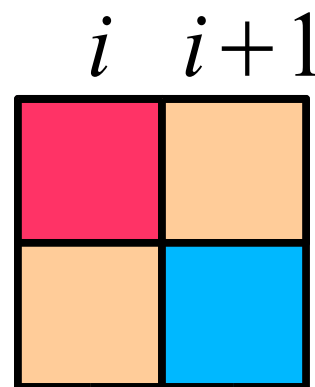
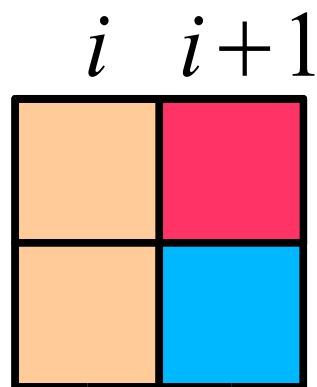
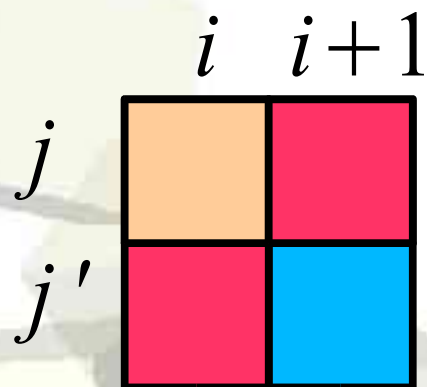
$$Y_{t+1}(j) \leftarrow Y_t(j) - 1$$



Case IV: active items j, j' at positions $i, i+1$



4 subcases when $j' \leq j$





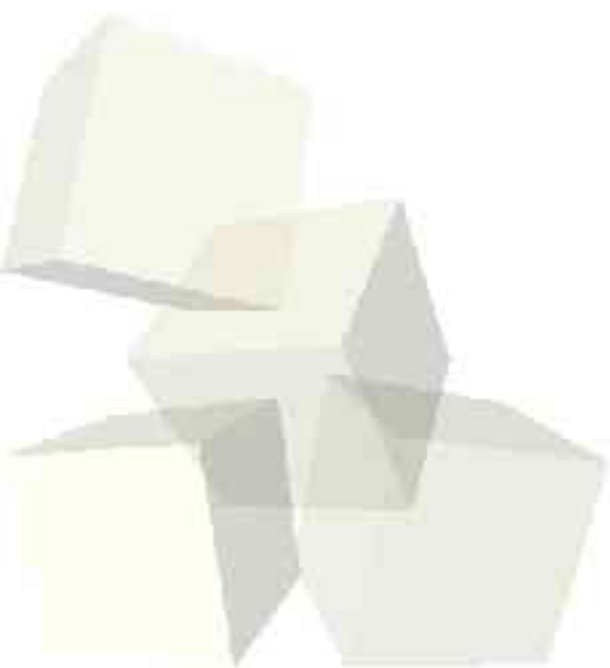
See what happens in each subcase

Action

If $B=0$ or $j' \leq j$ do nothing

else $Y_{t+1}(j) \leftarrow Y_t(j) - 1$

$Y_{t+1}(j') \leftarrow Y_t(j') + 1$





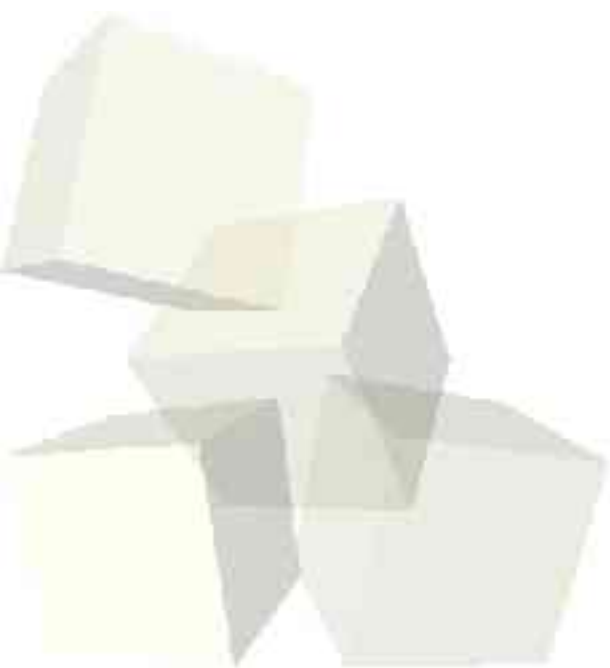
Time dependent

This update function is time dependent

The update depends on the state of the bounding chain

The bounding chain state depends on

$$U_t, U_{t-1}, U_{t-2}, \dots$$

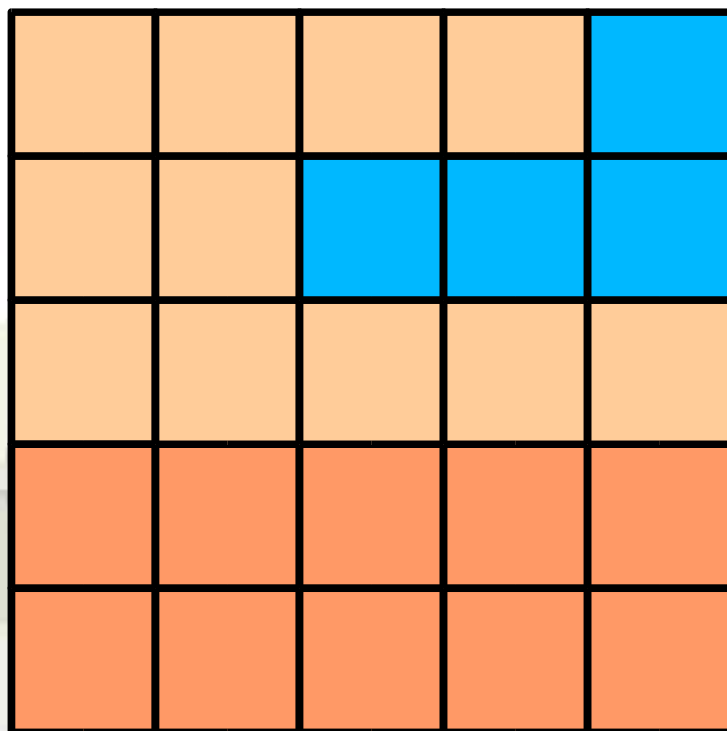




Key fact: The number of active states can only stay the same or increase

Call position i a hole if $Y_t(j) \neq \{1, \dots, i\}$ for all j

hole *hole*
↓ ↓

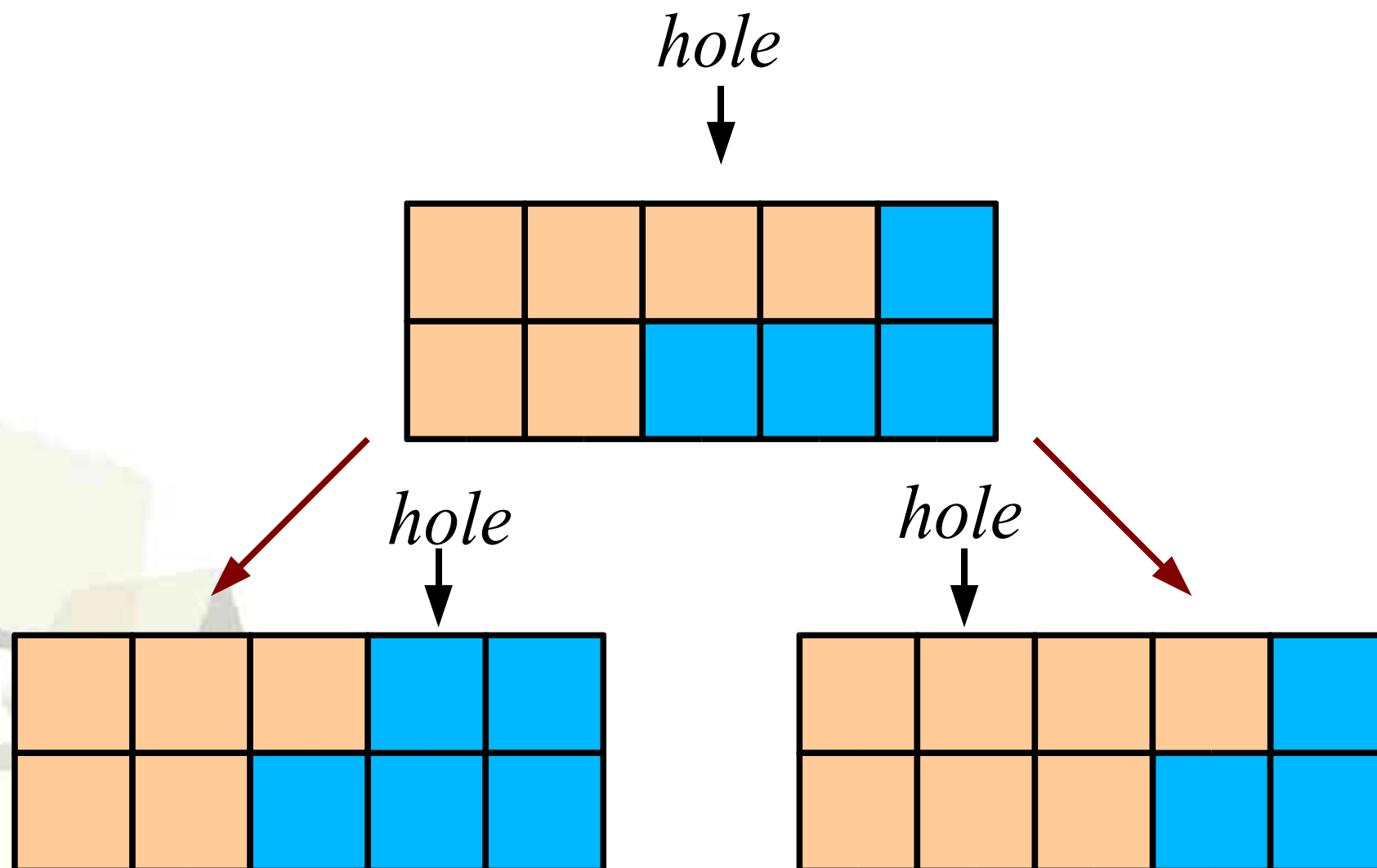


} *active items* $\{1, 2, 3\}$

item

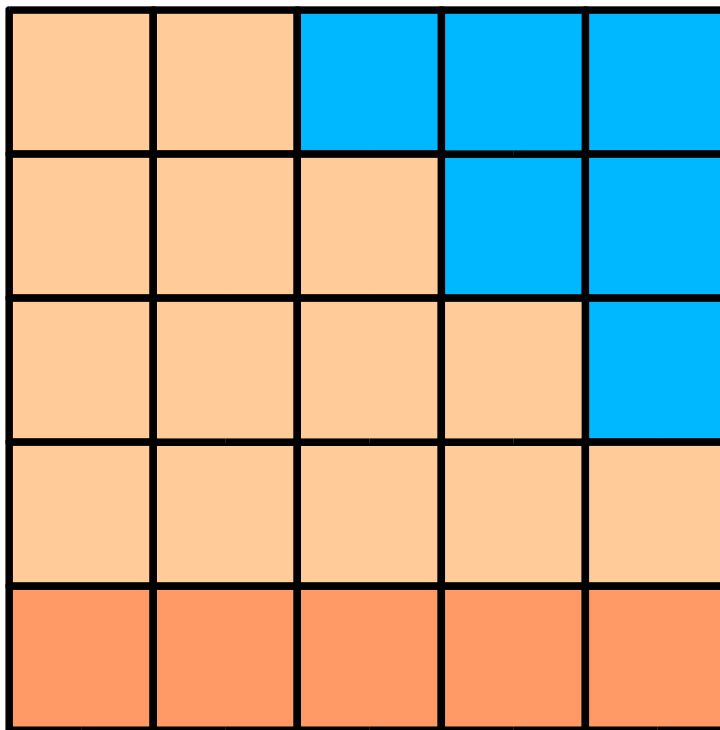


Roughly speaking: clear positions take a simple random walk on $\{1, \dots, n\}$





When a hole reaches n it creates a new active item



Time for algorithm is time for all holes to reach state n



Analysis continued

A hole needs n^2 moves in expectation to reach n

A hole is moved one in every $n/2$ steps

There are at most n holes at the start

Expected Running time

$$O(n^3 \ln(n))$$





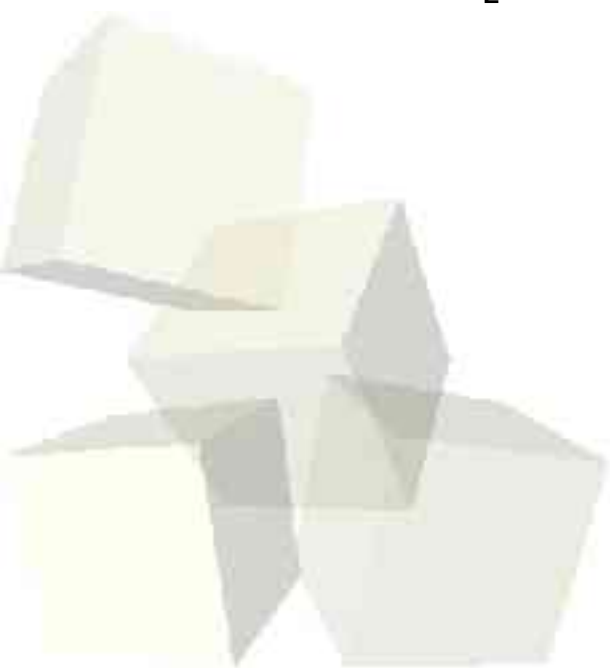
A potential proof

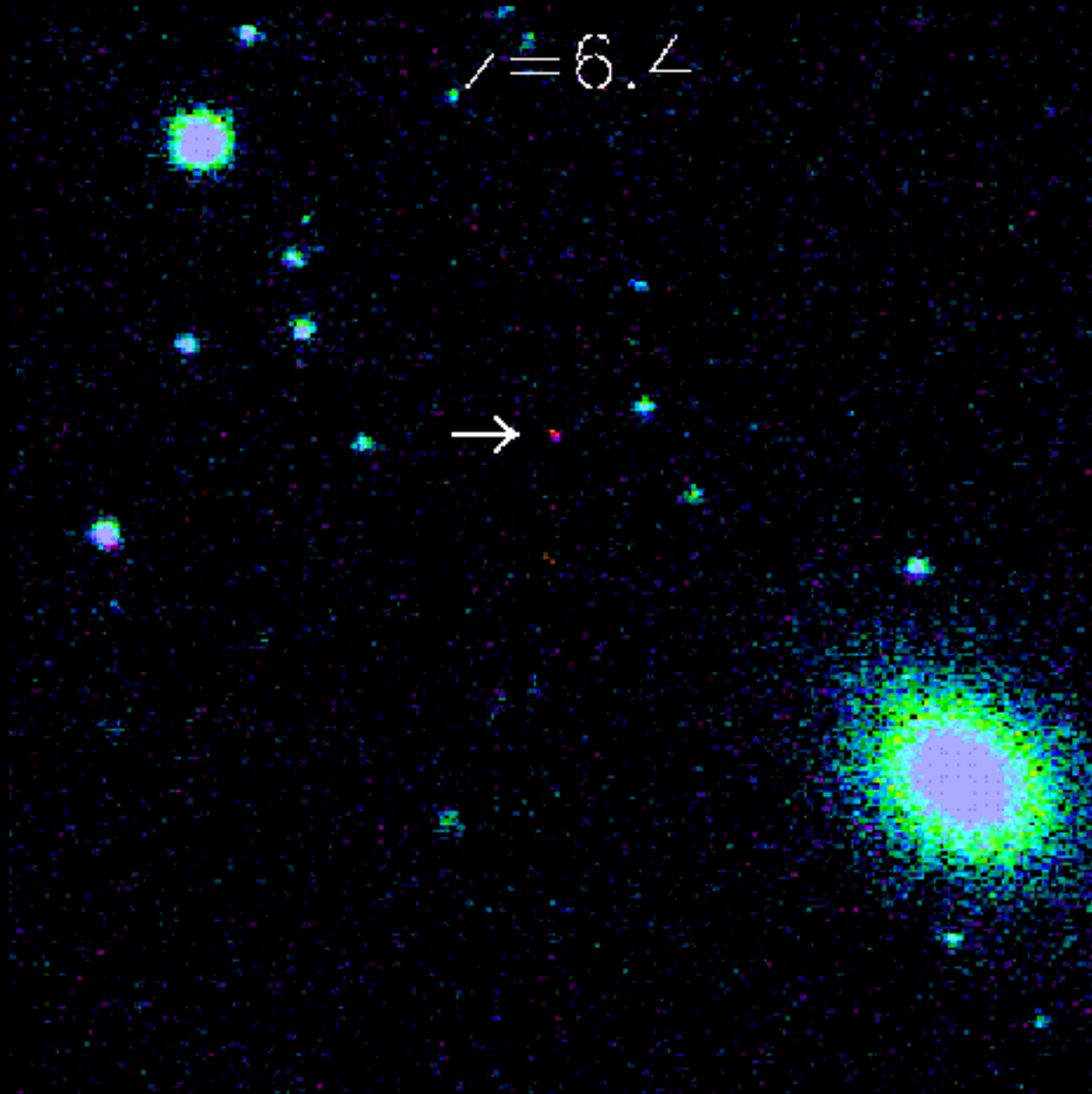
Create a potential function

$$\phi(Y_t) = \sum_{i \text{ is a hole}} [n^2 - i^2]$$

Can show

$$E[\phi(Y_{t+1}) | Y_t] \leq \phi(Y_t) \left[1 - \frac{2}{n^3} \right]$$



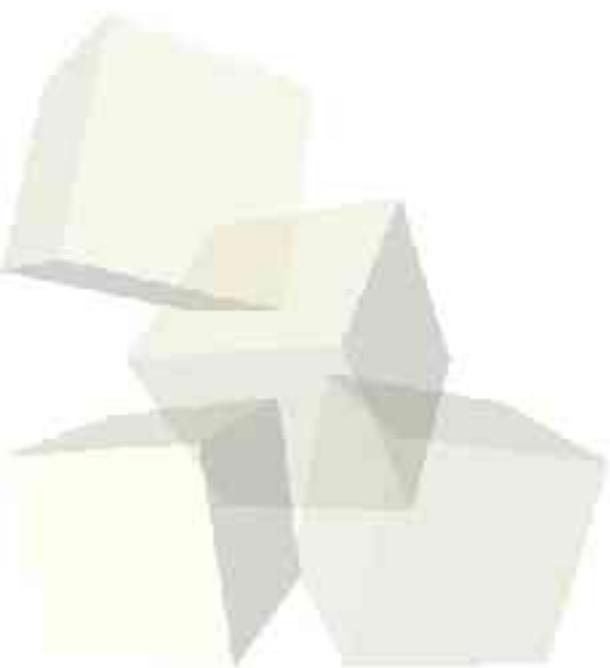
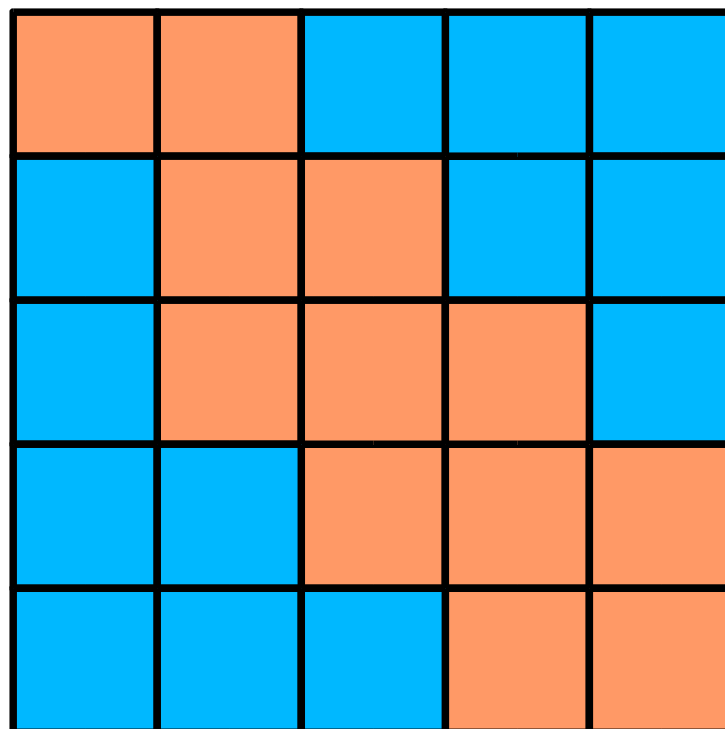


CREDIT: Sloan Digital Sky Survey at
Apache Point Observatory



A related problem

[Efron, Petrosian 1999] Quasar luminosity data is doubly truncated, making testing correlation difficult. They suggest nonparametric test: to find p -value, need samples from interval permutations

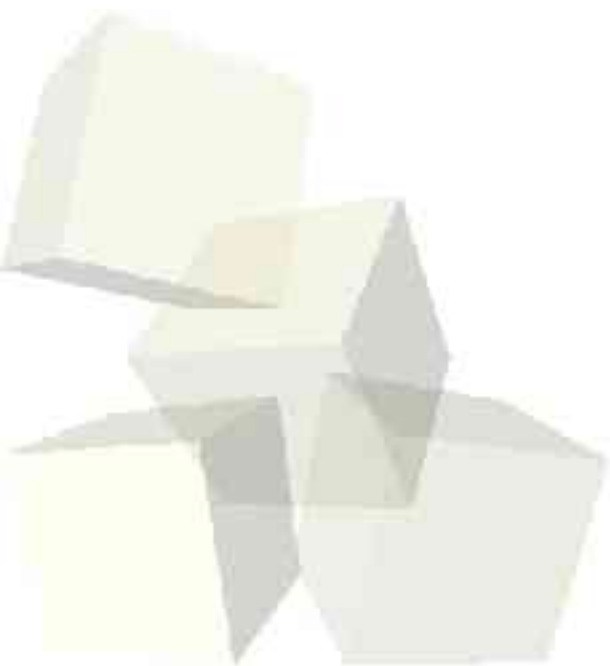




Analysis difficult

Unfortunately, holes cannot move freely in this example, and so currently no a priori estimates on the running time of this algorithm are known

On the other hand: who cares? Run the algorithm, if it's fast use it.





What are update functions?

Update functions are examples of couplings

A coupling runs two processes simultaneously

Marginally, each process looks like the original chain

Their moves can be dependent

A coupled process $\{A_t, B_t\}$ is faithful if

$$A_t = B_t \text{ implies } A_{t+1} = B_{t+1} \text{ for all } t$$

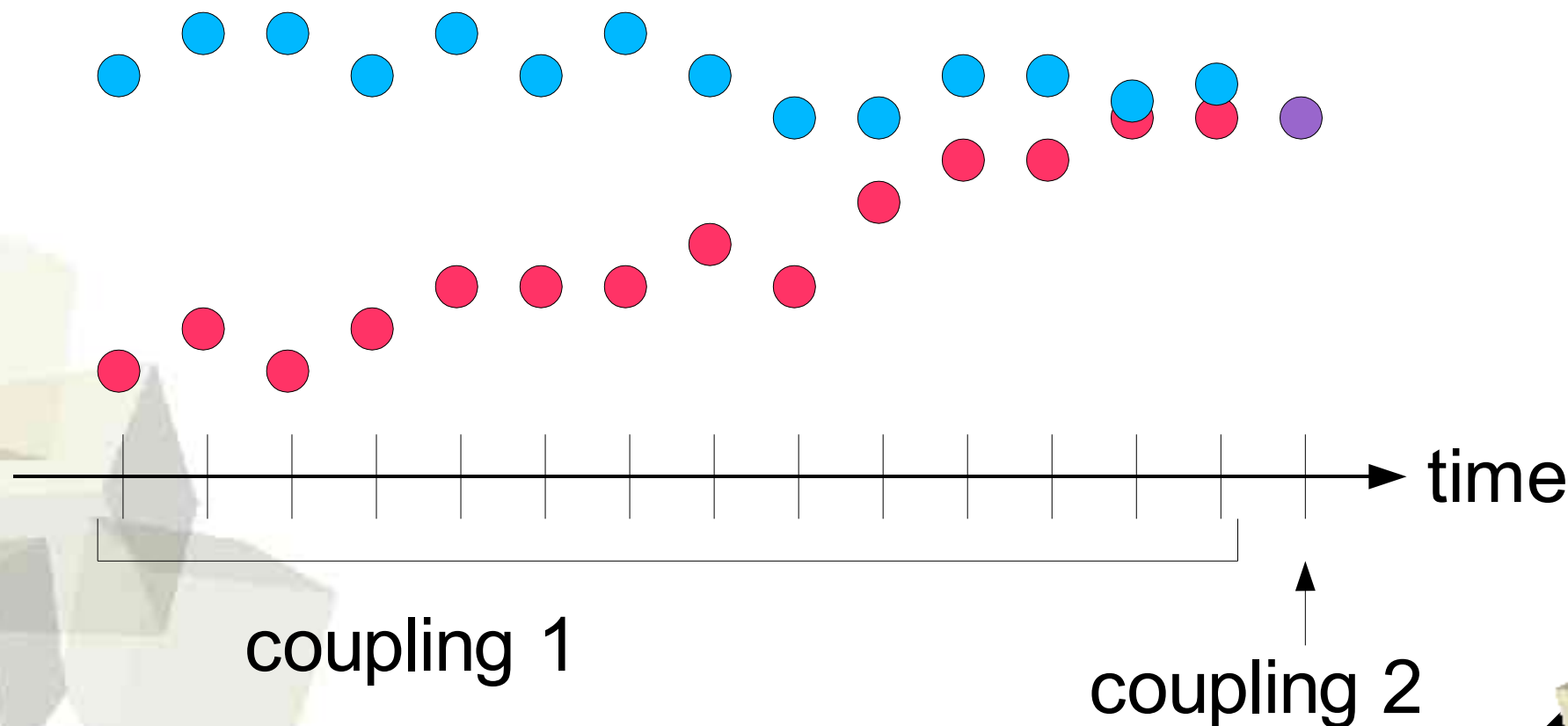
The coupling lemma says that the time until coupling is a bound on the mixing time of the Markov chain



Continuous problems

Time dependent couplings have already been used for continuous problems

Used for proving mixing times
Not for perfect sampling





Continuous Coupling

Type 1 coupling brings the states “close”

Type 2 coupling finishes the job

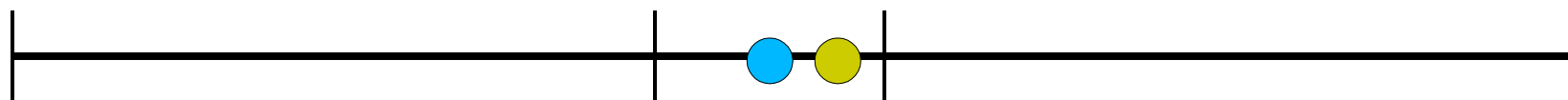
[Wilson 1999] Multishift coupling for type 2
works for unimodal distribution on moves

With time dependent CFTP, need much
weaker type 2 coupling





Continuous random walk on $[0, n]$



$$X_{t+1} \sim \text{Unif}[\max\{X_t - 1, 0\}, \min\{X_t + 1, n\}]$$

Type 1 update function

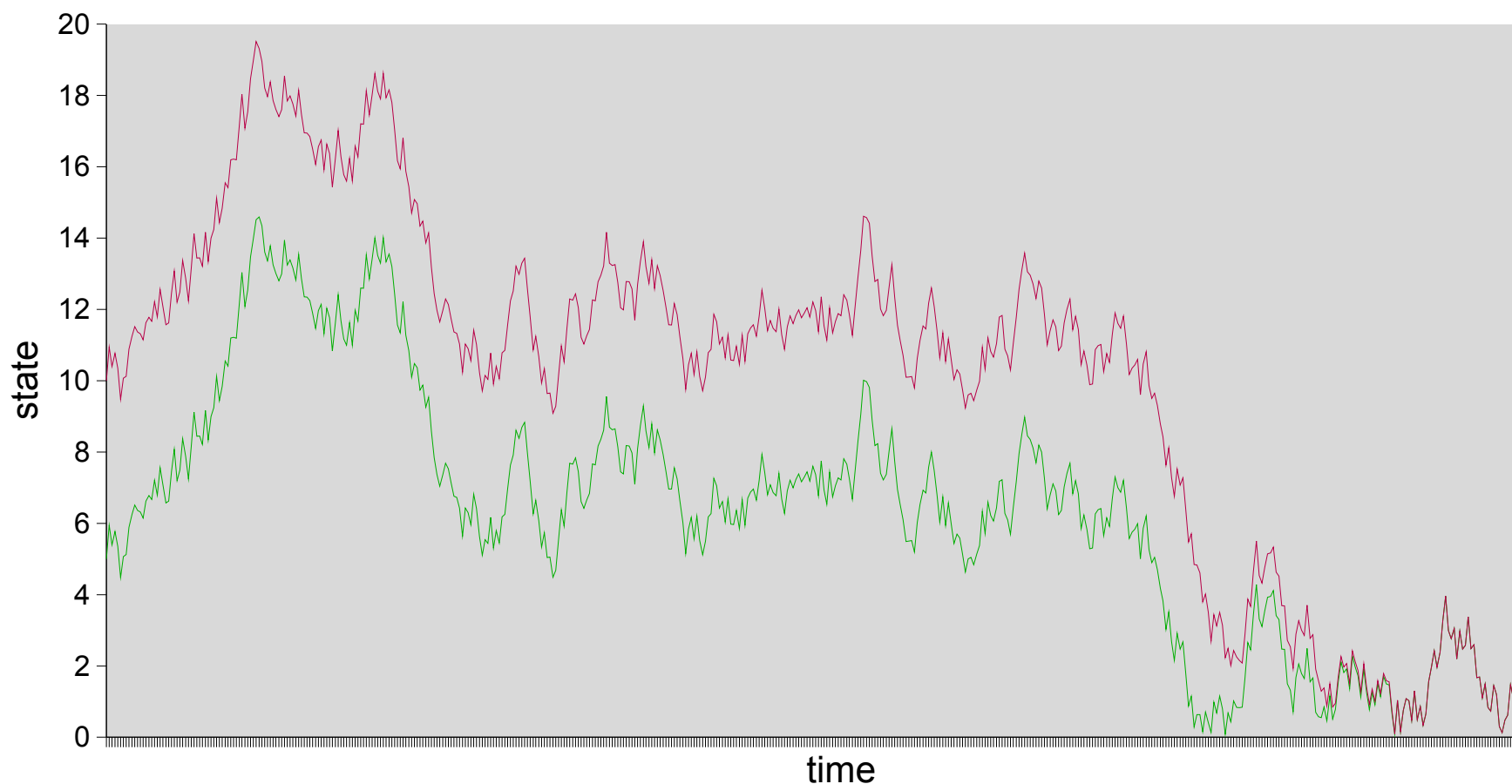
$$b = \min\{X_t + 1, n\}$$

$$a = \max\{X_t - 1, 0\}$$

$$f(x, u) = u(b - a) + a$$



Type 1 in action



Starting difference: 5 on $[0,20]$

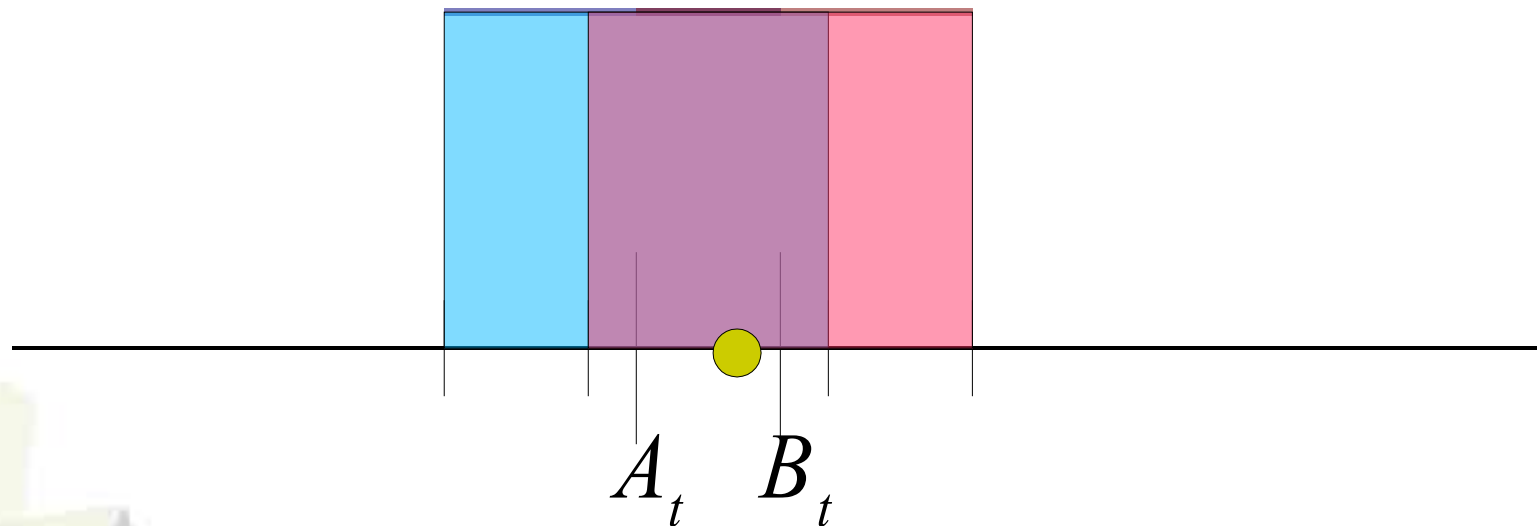
Number of steps: 500

Final difference: 0.000000078



Type 2 coupler

Idea: Distribution of A_{t+1}, B_{t+1} overlaps once A_t, B_t close



A draw that lands in shared density region works for both processes



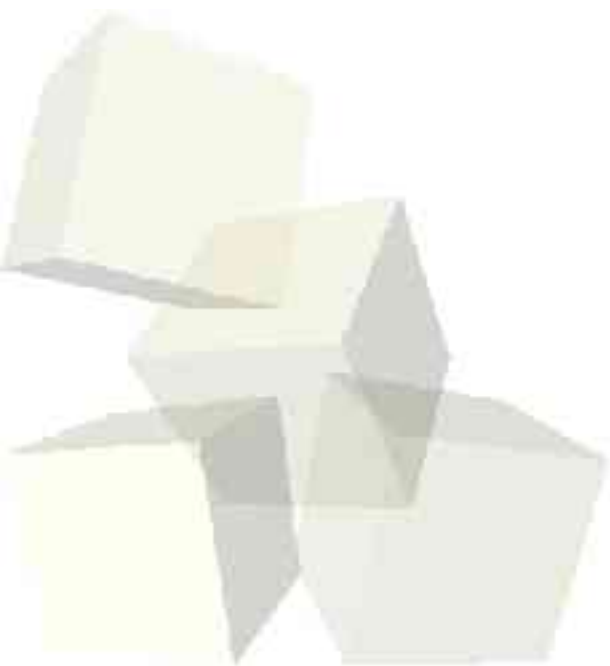
Multishift and time dependent

To perfectly sample

Use type 1 update function in time $[-T, 1]$

Use type 2 update function at time 1 to finish

For most problems, this idea makes
continuous no more difficult than discrete





Adding time dependent update functions increases the power of CFTP for perfect sampling

For linear extensions, provides fastest known method of generating samples

$$O(n^3 \ln(n))$$

Also works for interval permutations

Considerably simplifies algorithms for continuous problems



G. Brightwell and P. Winkler, Counting linear extensions. *Order* 8(3): 1—17, 1991

R. Bubley and M. Dyer, Faster random generation of linear extensions, *Disc. Math.* 201 (1999) 81—88

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S. Felsner and L. Wernisch. Markov chains for linear extension, the two-dimensional case. In *Proc. of 8th Annual ACM-SIAM Symposium on Discrete Algorithms*, (1997) 239—247

M. Huber. Perfect sampling using bounding chains. *Annals of Applied Probability*, 2004, to appear

M. Huber. Fast perfect sampling from linear extensions (submitted)



M. Jerrum, L. Valiant, and V. Vazirani, Random generation of combinatorial structures from a uniform distribution. Theoret. Comput. Sci. 43 (1986) 169—188

D. Wilson, Mixing times of lozenge tilings and card-shuffling Markov chains, Annals of Applied Probability, 2004

My website:

<http://www.math.duke.edu/~mhuber>

David Wilson's perfect sampling page

<http://dbwilson.com/exact/>